

Problem Setting

Let $(x^\mu, y^\mu) \in \mathbb{R}^d \times \mathcal{Y}$, $\mu \in [n]$, with $x^\mu \stackrel{iid}{\sim} \mathcal{N}(0_d, \Omega_0)$ and $y^\mu = f_*(x^\mu)$ a (random) target function. We consider a generalised linear estimation:

$$\hat{y} = \sigma \left(\frac{\theta^\top \varphi(x)}{\sqrt{k}} \right), \quad (1)$$

with *deep random features* (dRF):

$$\varphi(x) := \underbrace{(\varphi_L \circ \varphi_{L-1} \circ \dots \circ \varphi_2 \circ \varphi_1)}_L(x), \quad (2)$$

where the post-activations are given by:

$$\varphi_\ell(h) = \sigma_\ell \left(\frac{1}{\sqrt{k_{\ell-1}}} W_\ell \cdot h \right), \quad \ell \in [L]. \quad (3)$$

The entries of $\{W_\ell \in \mathbb{R}^{k_\ell \times k_{\ell-1}}\}_{\ell \in [L]}$ are $(W_\ell)_{ij} \stackrel{iid}{\sim} \mathcal{N}(0, \Delta_\ell)$.

Sample covariance matrices

Sample covariance matrix $\hat{\Sigma} := \mathcal{X} \mathcal{X}^\top / n \in \mathbb{R}^{d \times d}$ for $\mathcal{X} := (x_1, \dots, x_n)$, corresponding to the *population covariance matrix* Σ .

Gram matrix $\hat{\Sigma} := \mathcal{X}^\top \mathcal{X} / n \in \mathbb{R}^{n \times n}$ has the same non-zero eigenvalues. In the regime $d \sim n \gg 1$ the empirical spectral density $\mu(\hat{\Sigma}) := d^{-1} \sum_{\lambda \in \text{Spec}(\hat{\Sigma})} \delta_\lambda$ of $\hat{\Sigma}$ is approximately equal to the *free multiplicative convolution* of $\mu(\Sigma)$ and a Marchenko-Pastur distribution μ_{MP}^c with $c = d/n$,

$$\mu(\hat{\Sigma}) \approx \mu(\Sigma) \boxtimes \mu_{\text{MP}}^{d/n}. \quad (4)$$

The free multiplicative convolution $\mu \boxtimes \mu_{\text{MP}}^c$ is the unique distribution ν whose Stieltjes transform $m = m_\nu(z) := \int (x - z)^{-1} d\nu(x)$ satisfies the scalar *self-consistent equation*

$$zm = \frac{z}{1 - c - czm} m_\mu \left(\frac{z}{1 - c - czm} \right). \quad (5)$$

Gaussian universality of the test error

$\theta \in \mathbb{R}^k$ is obtained via the regularized *empirical risk minimization*:

$$\hat{\theta} = \underset{\theta \in \mathbb{R}^k}{\text{argmin}} \left[\sum_{\mu=1}^n \ell(y^\mu, \theta^\top \varphi(x^\mu)) + \frac{\lambda}{2} \|\theta\|^2 \right], \quad (6)$$

where $\ell: \mathcal{Y} \times \mathbb{R} \rightarrow \mathbb{R}_+$ is a convex loss function.

We assume that the labels are generated by a deep random neural network:

$$f_*(x^\mu) = \sigma^* \left(\frac{\theta_*^\top \varphi^*(x^\mu)}{\sqrt{k^*}} \right).$$

Here, $\theta_* \in \mathbb{R}^{k^*}$ and φ^* denotes composition $\varphi_{L^*}^* \circ \dots \circ \varphi_1^*$:

$$\varphi_\ell^*(x) = \sigma_\ell^* \left(\frac{1}{\sqrt{k_{\ell-1}^*}} W_\ell^* \cdot x \right).$$

The matched setting $\varphi = \varphi^*$ with the readout layer trained using a square loss corresponds to $\mathcal{Y} = \mathbb{R}$, $\ell(y, \hat{y}) = 1/2(y - \hat{y})^2$. Here (6) equals to

$$\hat{\theta} = 1/\sqrt{k} (\lambda I_k + 1/k X_L X_L^\top)^{-1} X_L y \quad (7)$$

where $X_L \in \mathbb{R}^{k \times n}$ contains last layer features column-wise and $y \in \mathbb{R}^n$.

Deterministic equivalents

The relationship (4) between the asymptotic spectra of Σ and $\hat{\Sigma}$, $\check{\Sigma}$ extends to eigenvectors, and the resolvents $\hat{G}(z) := (\hat{\Sigma} - z)^{-1}$, $\check{G}(z) := (\check{\Sigma} - z)^{-1}$ are asymptotically equal to *deterministic equivalents*.

Iteration over one layer

Consider a data matrix $X_0 \in \mathbb{R}^{d \times n}$ and $X_1 := \sigma_1(W_1 X_0 / \sqrt{d}) \in \mathbb{R}^{k_1 \times n}$. Furthermore, assume that the Gram matrix concentrates as

$$\left\| \frac{X_0^\top X_0}{d} - r_1 I \right\|_{\max} \prec \frac{1}{\sqrt{n}}, \quad \left\| \frac{X_0}{\sqrt{d}} \right\| \prec 1 \quad (8)$$

for some positive constant r_1 . For any deterministic A and Lipschitz-continuous σ_1 , for any $z \in \mathbb{C} \setminus \mathbb{R}_+$ (denoting $\langle A \rangle := \text{Tr} A / n$ for $A \in \mathbb{R}^{n \times n}$)

$$\left| \left\langle A \left[\left(\frac{X_1^\top X_1}{k_1} - z \right)^{-1} - \left(c_1(z) \frac{X_0^\top X_0}{d} + c_2(z) \right)^{-1} \right] \right\rangle \right| \prec \frac{\langle AA^* \rangle^{1/2}}{\sqrt{n}},$$

and similar result holds for the matrix $X_1 X_1^\top / k_1$. Furthermore, Assumption (8) holds true with X_0, r_1 replaced by X_1, r_2 , respectively.

Proof idea

The proof follows from the following sequence of approximations

$$\left(\frac{X_1^\top X_1}{k_1} - z \right)^{-1} \approx \left(c_0(z) \Sigma_X - z \right)^{-1} \approx \left(c_1(z) \frac{X_0^\top X_0}{d} + c_2(z) \right)^{-1}, \quad (9)$$

where $\Sigma_X := \mathbb{E}_{w \sim \mathcal{N}(0, I)} \sigma \left(\frac{X_0^\top w}{\sqrt{d}} \right) \sigma \left(\frac{w^\top X_0}{\sqrt{d}} \right) \in \mathbb{R}^{n \times n}$. The first approximation follows from [1], and the second one from Hermite series expansion.

The proposition can be iterated over arbitrary finite number of layers L

$$\left(\frac{X_L^\top X_L}{k_L} - z \right)^{-1} \approx \left(c'_1 \frac{X_{L-1}^\top X_{L-1}}{k_{L-1}} + c'_2 \right)^{-1} \approx \dots \approx \left(c_1 \frac{X_0^\top X_0}{d} + c_2 \right)^{-1}, \quad (10)$$

where c_1, c'_1, c_2, c'_2 are some functions of $z \in \mathbb{C} \setminus \mathbb{R}_+$.

In the proof, we assume fixed X_0 and random W_ℓ , leading to Σ_X . This approach facilitates iteration over the layers and appears in [2]. Another view is to consider $X_\ell X_\ell^\top / n$ as a sample covariance matrix with population covariance $\Omega_\ell := \mathbb{E}_{X_0} \frac{X_\ell X_\ell^\top}{n}$ since the matrix X_ℓ conditioned on W_1, \dots, W_ℓ has independent columns. The matrices are related as the population covariance and Gram matrices. We also derive a heuristic formula for Ω_ℓ .

Ridge universality of matched target

Let $\lambda > 0$. In the limit $n \sim d \sim k_\ell \gg 1$, under Assumption (8), the asymptotic test error of the ridge estimator (7) on the target (1) with $L = L^*$ and $\varphi_\ell^* = \varphi_\ell$ and additive $\mathcal{N}(0, \Delta)$ noise is given by:

$$\epsilon_g(\hat{\theta}) \xrightarrow{k \rightarrow \infty} \epsilon_g^* = \Delta (\langle \Omega_L \rangle \tilde{m}_L(-\lambda) + 1) - \lambda (\lambda - \Delta) \langle \Omega_L \rangle \partial_\lambda \tilde{m}_L(-\lambda)$$

where \tilde{m}_L can be recursively computed.

This implies Gaussian universality of this model, since (1) agrees with the asymptotic test error of data $x \sim \mathcal{N}(0_d, \Omega_L)$ derived in [3].

General case

The same result is shown to hold numerically for a much wider class of models, i.e., they belong to the *Gaussian universality class*. There is an equivalent Gaussian covariate model consisting of doing generalized linear estimation on $\check{\mathcal{D}} = \{v^\mu, \check{y}^\mu\}_{\mu \in [n]}$ with labels $\check{y}^\mu = f_*(1/\sqrt{k^*} \theta_*^\top u^\mu)$ and:

$$(u, v) \sim \mathcal{N} \left(\begin{array}{cc} \Psi_{L^*} & \Phi_{L^* L} \\ \Phi_{L^* L}^\top & \Omega_{L^*} \end{array} \right) \quad (11)$$

where $\Phi \in \mathbb{R}^{k^* \times k^*}$ and $\Psi \in \mathbb{R}^{k^* \times k^*}$ are the covariances between the model and target features and the target variance respectively. This provides an analogous contribution as [4] to the case of multi-layer random features.

Depth-induced implicit regularization

An insightful takeaway is that the activations in dRF (2) share the same population statistics as the activations in a deep *noisy* linear network

$$\varphi_\ell^{\text{lin}}(x) = \kappa_1^\ell \frac{W_\ell^\top x}{\sqrt{k_{\ell-1}}} + \kappa_*^\ell \xi_\ell, \quad (12)$$

where $\xi_\ell \sim \mathcal{N}(0_{k_\ell}, I_{k_\ell})$ is a Gaussian noise term.

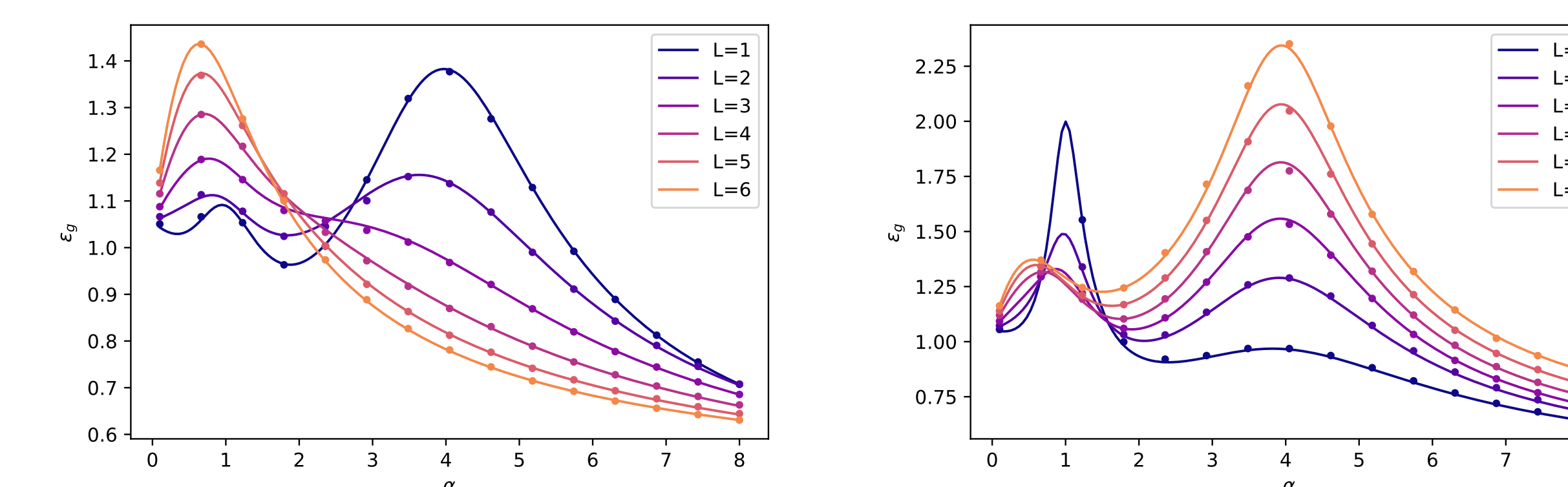


Figure 1: Learning curves for ridge regression on a 1-hidden layer target function ($\gamma_1^* = 2$, $\sigma_1^* = \text{sign}$) using a L -hidden layers learner with widths $\gamma_1 = \dots = \gamma_L = 4$ and $\sigma_{1, \dots, L} = \tanh$ activation (left) or $\sigma_{1, \dots, L}(x) = 1.1 \times \text{sign}(x) \times \min(2, |x|)$ clipped linear activation (right), for depths $1 \leq L \leq 6$. The regularization is $\lambda = 0.001$. *Solid lines* represent theoretical curves, while numerical simulations are indicated by *dots*. Two peaks, linear and non-linear, appear at $\alpha = n/d = 1$ and $\alpha = \gamma = 4$ respectively.

There exists an interplay between the two peaks, with higher noise ξ_L both helping to mitigate the linear peak, and aggravating the non-linear peak. The depth of the network plays a role in that it modulates the amplitudes of the signal part and the noise part.

References

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